

Sums of independent random variables

Recap: X, Y are indep. if $\forall A, B$

$$P(X \in A, Y \in B) = P(X \in A) \cdot P(Y \in B)$$

$$\Leftrightarrow f_{X,Y}(x,y) = f_X(x) \cdot f_Y(y) \text{ for all } x, y \in \mathbb{R}$$

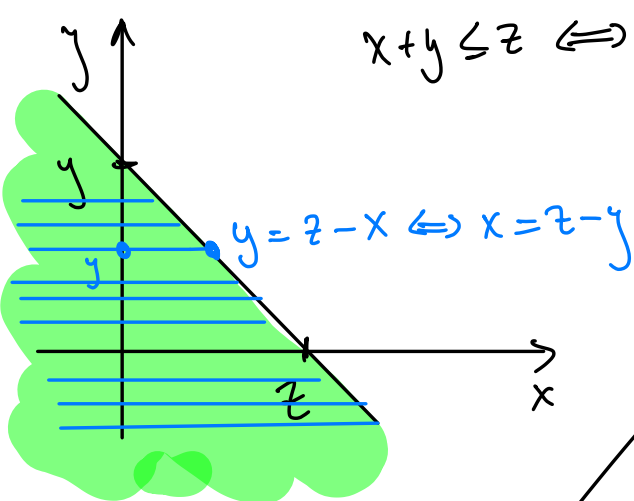
Q: Let X, Y be indep. random variables, and define $Z = X + Y$

Find the p.d.f. $f_Z(z)$ of Z .

A: Similar to strategy used for function of random var. ($g(X)$)

$$F_Z(z) = P(Z \leq z) = P(X + Y \leq z) = \iint_{x+y \leq z} \underbrace{f_{X,Y}(x,y)}_{f_X(x) f_Y(y)} dx dy$$

$$x + y \leq z \Leftrightarrow y \leq z - x \quad \leftarrow \text{given}$$



$$= \int_{-\infty}^{+\infty} \left(\int_{-\infty}^{z-y} f_X(x) f_Y(y) dx \right) dy$$

$$= \int_{-\infty}^{+\infty} f_Y(y) \left(\int_{-\infty}^{z-y} f_X(x) dx \right) dy$$

$$\underbrace{\left(\int_{-\infty}^{z-y} f_X(x) dx \right)}_{F_X(z-y)}$$

$$= \int_{-\infty}^{+\infty} F_X(z-y) f_Y(y) dy$$

Differentiate w.r.t z :

$$f_Z(z) = \frac{d}{dz} F_Z(z) = \frac{d}{dz} \int_{-\infty}^{+\infty} F_X(z-y) f_Y(y) dy =$$

under good enough hypothesis

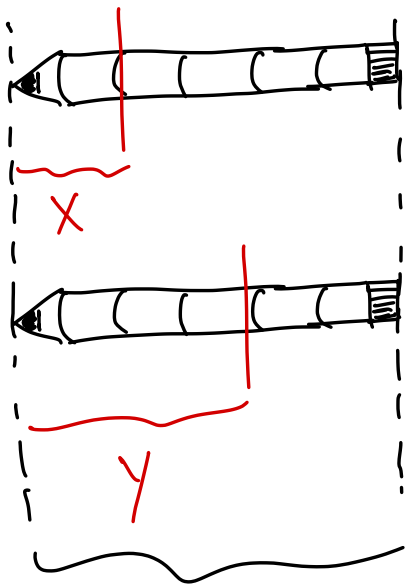
$$= \int_{-\infty}^{+\infty} f_Y(y) \underbrace{\frac{d}{dz} F_X(z-y)}_{f_X(z-y) \cdot 1} dy = \int_{-\infty}^{+\infty} f_X(z-y) f_Y(y) dy$$

Upsshot: $f_Z(z) = f_{X+Y}(z) = \int_{-\infty}^{+\infty} f_X(z-y) f_Y(y) dy = (f_X * f_Y)(z).$

"Convolution product of f_X and f_Y "

Example: Suppose $X, Y \sim \text{Uniform}([0,1])$. Find the p.d.f. of $X+Y$.

(Suppose two wooden pencils are broken into 2 parts at a random uniformly distributed break point. What is the probability that the sum of lengths of the two tip parts is at most $2/3$ of the length of original pencil?)



$$X, Y \sim \text{Uniform}([0,1]) \Rightarrow f_X(x) = \begin{cases} 1 & \text{if } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

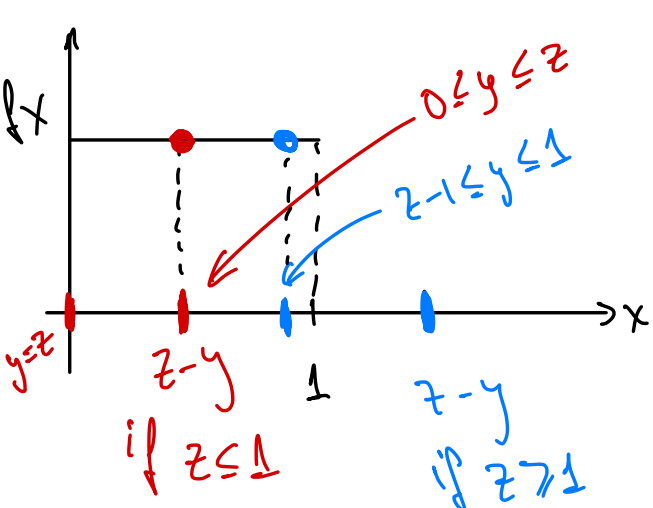
$$f_Y(y) = \begin{cases} 1 & \text{if } 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

$Z = X+Y$ takes values on $[0,2]$.

$$f_{X+Y}(z) = (f_X * f_Y)(z) = \int_{-\infty}^{\infty} f_X(z-y) f_Y(y) dy$$

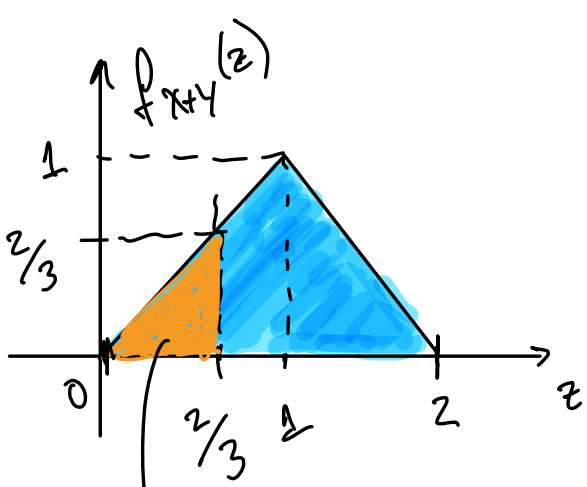
$\begin{matrix} \text{|||} \\ 1 \text{ if } y \in [0,1] \end{matrix}$

$$= \int_0^1 f_X(z-y) dy = \begin{cases} \int_0^z 1 dy = z & \text{if } 0 \leq z \leq 1 \\ \int_{z-1}^1 1 dy = 1 - (z-1) = 2-z & \text{if } 1 \leq z \leq 2 \\ 0 & \text{otherwise} \end{cases}$$



Upside:

$$f_{X+Y}(z) = \begin{cases} z & \text{if } 0 \leq z \leq 1 \\ 2-z & \text{if } 1 \leq z \leq 2 \\ 0 & \text{otherwise} \end{cases}$$



The prob. that the sum of the tip parts of broken pencils is $\leq \frac{2}{3}$ of original length of a pencil is:

$$\rightarrow P(X+Y \leq \frac{2}{3}) = \frac{1}{2} \left(\frac{2}{3}\right)^2 = \boxed{\frac{2}{9}}$$

Also, note that:



Sums of uniform random variables is not a uniform random variable!

How are sums of (independent) random variables distributed?

$$\begin{aligned} X &\sim \text{Normal}(\mu_1, \sigma_1) \\ Y &\sim \text{Normal}(\mu_2, \sigma_2) \end{aligned} \implies X+Y \sim \text{Normal}\left(\mu_1+\mu_2, \sqrt{\sigma_1^2+\sigma_2^2}\right)$$

$$\begin{aligned} X &\sim \text{Poisson}(\lambda_1) \\ Y &\sim \text{Poisson}(\lambda_2) \end{aligned} \implies X+Y \sim \text{Poisson}(\lambda_1+\lambda_2)$$

but: $X, Y \sim \text{Uniform}$ but $X+Y \not\sim \text{Uniform}$

$X, Y \sim \text{Exponential}$ but $X+Y \not\sim \text{Exponential}$!

Expected value of functions of random variables

Recall: $E(g(X)) = \int_{-\infty}^{+\infty} g(x) f_X(x) dx$

Def: If X, Y are cont. random variables, and $g: \mathbb{R}^2 \rightarrow \mathbb{R}$,
then

$$E(g(X, Y)) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} g(x, y) f_{X, Y}(x, y) dx dy$$

If X, Y are discrete, then

Joint prob. mass function
 $p(x, y) = P(X=x, Y=y)$.

$$E(g(X, Y)) = \sum_x \sum_y g(x, y) p(x, y)$$

In Lecture 10 (videos), we proved $E(\cdot)$ is linear, i.e.

$$E(aX + bY) = aE(X) + bE(Y)$$

if X, Y are discrete. Now, let us extend this to cont. r.v.

Prop.: $E(\cdot)$ is linear for cont. random variables.

Pf.: X, Y cont. random variables, $a, b \in \mathbb{R}$. We want to

show $E(aX + bY) = aE(X) + bE(Y)$.

$$E(aX + bY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (ax + by) \cdot f_{X, Y}(x, y) dx dy =$$

$$g(x, y) = ax + by$$

$$= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} ax f_{X, Y}(x, y) + by f_{X, Y}(x, y) dx dy$$

$$= a \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x f_{X, Y}(x, y) dx dy + b \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f_{X, Y}(x, y) dx dy$$

$$= a \int_{-\infty}^{+\infty} x \left(\int_{-\infty}^{+\infty} f_{X,Y}(x,y) dy \right) dx + b \int_{-\infty}^{+\infty} y \left(\int_{-\infty}^{+\infty} f_{X,Y}(x,y) dx \right) dy$$

$f_X(x)$
 $f_Y(y)$

$$= a \int_{-\infty}^{+\infty} x f_X(x) dx + b \int_{-\infty}^{+\infty} y f_Y(y) dy = a E(X) + b E(Y).$$

$E(X)$
 $E(Y)$
□

Covariance: $\text{Cov}(X, Y) = E((X - E(X))(Y - E(Y)))$

$$= E(XY) - E(X)E(Y).$$

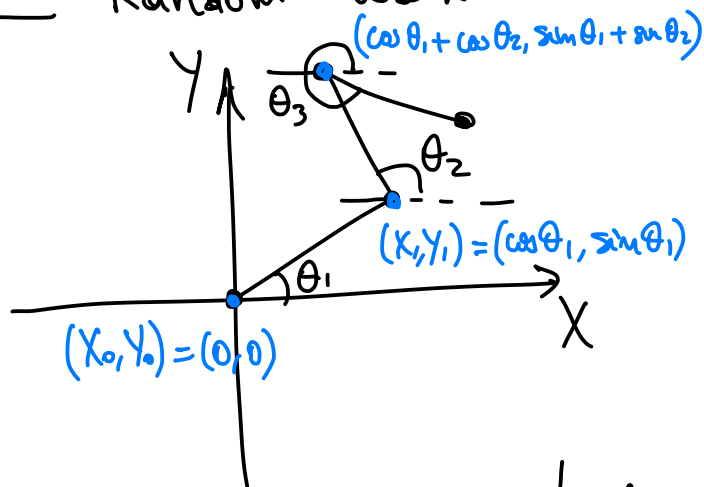
The above also holds in the same way for discrete and cont. random variables.

Cor: If X, Y are independent rand. var., then

$$\text{Cov}(X, Y) = 0 \iff E(XY) = E(X)E(Y)$$

Also, $E(g(X)h(Y)) = E(g(X))E(h(Y))$

Ex: Random walk in the plane, n steps:



Take steps of length 1 in a random and independent direction:

$$\theta_i \sim \text{Unif}([0, 2\pi])$$

at each step $1 \leq i \leq n$.

Q: Find the expected value of the square of the distance to the origin after n steps.

Position after n steps: $\left(\sum_{i=1}^n \cos \theta_i, \sum_{i=1}^n \sin \theta_i \right)$

$$E \left(\left(\sum_{i=1}^n \cos \theta_i \right)^2 + \left(\sum_{i=1}^n \sin \theta_i \right)^2 \right) \stackrel{\text{linearity}}{=} \dots$$

$$= E \left(\left(\sum_{i=1}^n \cos \theta_i \right)^2 \right) + E \left(\left(\sum_{i=1}^n \sin \theta_i \right)^2 \right)$$

$$= E \left(\sum_{i=1}^n \cos^2 \theta_i + \sum_{i \neq j} \cos \theta_i \cos \theta_j \right)$$

$$+ E \left(\sum_{i=1}^n \sin^2 \theta_i + \sum_{i \neq j} \sin \theta_i \sin \theta_j \right)$$

linearity

$$\underline{=} \sum_{i=1}^n E(\cos^2 \theta_i) + \sum_{i \neq j} E(\cos \theta_i \cos \theta_j)$$

$$+ \sum_{i=1}^n E(\sin^2 \theta_i) + \sum_{i \neq j} E(\sin \theta_i \sin \theta_j)$$

$$= \sum_{i=1}^n \left(E(\cos^2 \theta_i) + E(\sin^2 \theta_i) \right) + \sum_{i \neq j} E(\cos \theta_i \cos \theta_j) + E(\sin \theta_i \sin \theta_j)$$

$$= \sum_{i=1}^n E(\underbrace{\cos^2 \theta_i + \sin^2 \theta_i}_1) + \sum_{i \neq j} E(\cos \theta_i) E(\cos \theta_j) + E(\sin \theta_i) E(\sin \theta_j)$$

$$\underline{=} n + 0$$

$$= n.$$

$$\left[\begin{array}{l} E(\cos \theta_i \cos \theta_j) = E(\cos \theta_i) E(\cos \theta_j) \text{ by } \underline{\underline{\text{indep.}}} \\ E(\sin \theta_i \sin \theta_j) = E(\sin \theta_i) E(\sin \theta_j) \end{array} \right.$$

$$\left[\begin{array}{l} E(\cos \theta_i) = \frac{1}{2\pi} \int_0^{2\pi} \cos \theta_i d\theta_i = 0 \\ E(\sin \theta_i) = \frac{1}{2\pi} \int_0^{2\pi} \sin \theta_i d\theta_i = 0 \end{array} \right.$$

$$\underline{A} = E(\text{square dist. to origin}) = \# \text{ steps} = n.$$